

How Intangible Assets Affect the Corporate Financial Performances and How It Varies from Sector – to – Sector?

**Geoffrey VanderPal
Purdue University Global**

To explore the varied perspectives of intangible assets and corporate financial performance nexus, this study employs various measures, i.e. Generalized Method of Moments (GMM). Analysis reveals significant variances in different asset classes and in different sectors. The findings provide insights in risk-return paradigm of intangible investments and the successive returns besides helping the policy makers to settle the priority sector to get the expected result in line with the country's investment policy.

Keywords: Intangible Assets, Corporate Finance, Financial Performance, investment policy, sector performance

INTRODUCTION

Within an industry, intangible asset(s) is the main factor of differentiation and competitiveness of enterprises (Penman, 2009). Hence, in order to gain competitive advantage, firms undertake costly activities to develop innovations (Thatcher & Pingry, 2009). Recognizing the significance, in recent years, there has been an increasing interest among academics from different field of studies to understand the relationship dynamics between intangible assets and financial performance. At large, intangibles have positive effects on firm value and profitability, and an indicator for future financial performance (Chen, Cheng, & Hwang, 2005). However, previously published studies on the effect of intangible assets are not consistent. What is less clear is the nature of relationships for diverse sectors – whether and how such relationship dynamics varies. This paper attempts to show sectoral differences of the intangible assets - financial performance nexus. Besides, the Generalized Method of Moments (GMM) are employed to catch on the lag effect bearing in mind that the intangible expense outcome is not immediate.

This study aims to contribute to this growing area of research by exploring the relationship from diverse viewpoints. The study aims to examine the role of corporate intangible assets on their financial performance. Besides looking into the intangible expense, it considers the ratio of intangible assets to firm's total assets, to recognize how much the possession of intangibles are beneficial from the perspective of firm's asset holding.

The next section presents the literature review on intangibles and present the gap in the existing literature. This is followed by a section that describes the data, methodology and the econometric aspects of the study and present the empirical models to achieve the study objective. Then, the main findings and discussion are presented. The paper concludes with a discussion on the implications of the findings.

LITERATURE REVIEW

The current study reflects a knowledge-based perspective of firms which consider organizations as repositories of knowledge (Grant, 1996; Kogut & Zander, 1992). This view highlights the role of intangible resources, in combination with physical assets, in developing the required competitive advantage (Denicolai, Zucchella, & Strange, 2014). Such an advantage is viable if firms can transmit their investments in creating intangible assets that would enhance their financial performance. There are numerous theoretical perspectives that describe the relationship, which include the dynamic resource-based view (Helfat and Peteraf, 2003), and this has been confirmed by a number of empirical studies (e.g. Eberhart, Maxwell & Siddique, 2004; Lome, Heggeseth & Moen, 2016).

The study of competitive advantage driven by Intellectual property (IP) requires the identification of certain critical factors, such as those referred to by the literature here. These elements allow businesses to enjoy competitive advantages in terms of market power. The expenditures signal the strategic positioning of a firm and significantly put a strain on the firm's financial performances (Lantza & Sahutb, 2005; Rivette & Klein, 2008). Investors also accept investments in intangibles without any reserves (Garcia-Garcia & Magdaleno, 2010). Firms can therefore enjoy extra benefits by such investments in creating intangibles.

Since businesses are becoming more complex, dynamic and globally competitive, knowledgeable and intelligent workers have become valuable intangible assets who contribute towards creating value. Hence, it has been extensively established that at present, the most essential value generating resources are intangible in nature, and to be precise, are related to the skills and knowledge embedded in an organisation (Kiant, Andreeva, & Pavlov, 2013). Innovation is usually positively correlated with return on assets (Sher and Yang, 2005; Gamayuni, 2015). Yet, the propensity to invest in intangible assets are not homogeneous and surges according to the firm's size, human capital, and historical intangible asset base (Arrighetti, Landini, & Lasagni, 2014). Moreover, the value of intangible assets is more volatile than the value of tangible assets, and any changes increase the difference between the book value and market value (Garger, 2010).

Nonetheless, for US firms, the value of total assets increased by 57% when intangible capital are measured in addition to conventional financial accounts (Hulten & Hao, 2008). There is a positive influence by internal intangible capital on firms' productivity levels (Marrocu, Paci, & Pontis, 2012). The ownership of intangible assets also translates into superior organisational performance (Menor, Kristal, & Rosenzweig, 2007; Hsu & Sabherwal, 2011). Moreover, firms with feasible intangible assets are able to not only perform and grow locally but also internationally. Considering the firms' knowledge-based views, more knowledge assets have a positive impact on foreign sales intensity; which in turn influences firms' financial performances altogether (Denicolai et al., 2014). Analyses largely indicate that investment in intangibles enhances the value of a company (VanderPal, 2015). However, Denicolai et al., (2014) argue that such progressive impact of intangible assets can be enjoyed only up to a certain point. Their analysis of inverse quadratic relationship suggested the need to balance knowledge assets with complementary assets in order to achieve a higher degree of international performance.

Past studies have discussed the impacts of possessing intangible assets and consequential performance of firms, particularly their financial performance. The effect of such expenditure was found heterogeneous for growing or shrinking firms (Coad & Rao, 2009). The inconclusiveness in the current literature points to the need for further investigation to ascertain the impact of intangibles assets on financial performance.

METHODOLOGY AND ANALYSIS

Different methodologies employed to accomplish the study objectives where the key interest is to find out the how firms' financial performances are affected by investing intangibles. Sectoral segregation is made to comprehend the sector-wise idiosyncrasy. For this study, we collected the data from Compustat's S&P 500 companies. Data frequency is annual and it covers a range from 1979-2015.

In statistics, ordinary least squares (OLS) is a method for estimating the unknown parameters in a linear regression model. The OLS is considered a classical estimation method as the OLS estimator provides minimum-variance mean-unbiased estimation when the errors are homoscedastic and serially uncorrelated. However, one important issue that most models have to concern is the endogeneity problem which occurs when an explanatory variable correlates with the error term due to omitted variables, measurement errors, or simultaneity (Wooldridge, 2006). For the current study, companies' ownership of intangible assets are endogenous and is correlated with the error term and the classical OLS regression model might produce inefficient regression coefficient. To overcome these possible problems, the study uses the generalized method of moments (GMM) estimators developed for dynamic panel data that was first introduced by Hansen (1982) and proposed by Holtz-Eakin et al. (1988) and Arellano & Bond (1991). The GMM estimator has several advantages in particular for this study. Firstly, the GMM is an appropriate method for the research data structure and it performs well for the unbalanced dataset. Secondly, it can reduce the endogeneity problem due to the potential correlation between regressors and the error term. Thirdly, this research uses lagged dependent variables, thus, the dynamic GMM panel is the most appropriate method to address this type of data structure. Fourth, dynamic GMM panel data estimation is more appropriate in cases where some unobservable factors affect both the dependent variable and the explanatory variables and some explanatory variables are strongly related to past values of the dependent variable. Furthermore, introducing lagged values of the dependent variable in OLS estimators may seriously bias estimated coefficients (Nickell, 1981). In consideration of the above, heteroskedasticity and the properties of our panel dataset, Arellano and Bond's two-step difference GMM estimator is used. This dynamic Generalised Methods of Moments (GMM) estimator ensures a consistent and reliable estimation of the parameters of interest (Roodman, 2006). In general, the consistency of GMM estimator depends on the validity of the assumption that the error terms do not exhibit serial correlation and on the validity (exogeneity) of its instruments. To validate these assumptions, STATA¹ offers two sets of specification tests. The first set constitutes Sargan² and Hansen test³ of over-identification.

To check for first-order serial correlation in levels, we look for second-order correlation in differences AR (2) (Mileva, Bruhn, & Weickert, 2007). Autocorrelation in levels indicates that lags of the dependent variable (and any other variables used as instruments) are not strictly exogenous but in fact endogenous, thus bad instruments. Failure to reject the null hypotheses of the over-identification and serial correlation tests gives support to our model.

To eliminate the potential bias caused by omitted heterogeneity, we can either use fixed effects or random effects models. If the independent variables are uncorrelated with the unobserved effect (μ_i), the fixed effects estimator is consistent but inefficient, whereas the random effects estimator is consistent and efficient. If the independent variables are correlated with the unobserved heterogeneity (μ_i), the fixed effects estimator is consistent, while the random effects estimator is inconsistent (Baum, 2006). So, to identify the appropriate estimation model, we run the Hausman test. If the null hypothesis is rejected, then we conclude that μ_i is correlated with the independent variables, i.e. the fixed effects is the appropriate method (Pasiouras & Kosmidou, 2007; Petria, Capraru, & Ihnatov, 2015). Moreover, assuming homoscedasticity of error terms in the presence of heteroscedasticity, as well as having autocorrelated disturbances, produces consistent but inefficient estimates, and the standard errors of these estimates will be biased (Baltagi, 2005). Therefore, we shall estimate robust standard errors to correct for the possible presence of these issues.

At this level, we empirically test the effect of intangible assets on firm's financial performance using this model.

$$FIN_{i,t} = \alpha_i + \tau_t + \beta_1 INT_{it} + \beta_2 Cit + \epsilon_{it} \quad (1)$$

where: FIN_{it} represents financial performance; α_i is a dummy variable to control for time-invariant firm-specific factors (e.g., assets, sales etc.); τ_t is a year dummy variable to control for time-varying common shocks; INT_{it} is for intangible assets, Cit is a vector of firm-specific variables; ϵ_{it} is the normally distributed error term; i and t are indexes for cross-section and time-series, respectively, with $i = 1, \dots, N$,

$t = 1, \dots, T$. Positive sign of β_1 and significant p value indicates the existence of mission drift. Equation (1) specifies the overall model for the study. The equation is used to define different sectoral models by including all the variables and then modifying it for different sectors' data.

TABLE 1
LIST OF VARIABLES

Variable Name	Symbol
Assets-Total	AST
EBIT	EBI
Intangibles-Other	INT
Net Income (Loss)	NEI
Operating Income Before Depreciation	OPI
Price-Close Calendar Year	PRI
ROA	ROA
ROE	ROE
Sales-Net	SAL
Intangible Assets Ratio	INT / AST

Following the methodology explained, we report and analyse the various estimations (i.e. GMM and static panel threshold). In general, the results of most estimations indicate the significance of the lag dependent variable (i.e.), in line with earlier findings in other empirical studies (Name some references) that what confirm the appropriateness of using the GMM and Panel threshold technique.

Descriptive Statistics

Table 2 summarizes the median values of our main variables of corporate firms. It presents the summary statistics for the aggregated data. Clearly, the data are characterised by their heterogeneity, where the differences among corporate firms are significant.

TABLE 2
DESCRIPTIVE STATISTICS

Variable	No. of Observations	Mean	Standard Deviation	Minimum	Maximum
ROA	14314	5.259422	11.52165	-577.85	90.66
ROE	14008	15.35914	159.9642	-14132	7038.46
INT	6484	1966.882	7134.023	0	169054
INTAST	14322	0.028055	0.0720645	0	0.810795
LAST	14322	8.569283	1.930882	-1.17766	14.76063
LNEI	13037	5.611215	1.773562	-6.90776	11.56001
LEBI	13714	6.209501	1.721655	-3.07911	11.17367
NEI	14314	836.5912	2946.255	-99289	104821
OPI	13723	2238.955	5350.579	-76735	81730
PRI	13913	32.88359	59.03339	0.01	1971.25

Correlation Coefficient

Table 3 provides the matrix of Pearson correlation coefficients that based on the results, indicates relatively weak association between the variables.

**TABLE 3
CORRELATION COEFFICIENT**

Variable	ROA	ROE	INT	INTAST	LAST	LNEI	LEBI	NEI	OPI	PRI
ROA	1									
ROE	0.1457	1								
INT	-0.1131	-0.0134	1							
INTAST	-0.1261	-0.0032	0.4984	1						
LAST	-0.1836	-0.0041	0.446	0.1504	1					
LNEI	0.2576	0.0504	0.3695	0.0874	0.8619	1				
LEBI	0.1405	0.0516	0.3917	0.1272	0.9109	0.9449	1			
NEI	0.1852	0.0309	0.4036	0.062	0.6526	0.7036	0.6845	1		
OPI	0.052	0.0173	0.5119	0.0731	0.7103	0.681	0.7075	0.9282	1	
PRI	0.1306	0.0221	0.0422	0.0651	0.1722	0.2282	0.222	0.1507	0.1146	1

**TABLE 4
BASELINE MODEL: FINANCIAL PERFORMANCE AND INTANGIBLE ASSETS**

	ROA	ROE	ROA	ROE
	INT	INT	INT/AST	INT/AST
L.ROA	-0.0382*** (-8.13)		0.00416 (1.15)	
L.ROE		0.0357** (2.97)		-0.0286*** (-4.06)
INT	0.00000637 (0.38)	- (-0.75)		
INTAST			-2.687*** (-4.12)	-46.60 (-1.62)
LAST	-5.610*** (-43.81)	-9.908** (-3.11)	-5.414*** (-78.21)	-10.62*** (-3.47)
LNEI	2.750*** (31.70)	5.148* (2.57)	2.327*** (46.21)	5.419* (2.45)
LEBI	2.034*** (15.15)	4.803 (1.55)	2.413*** (29.58)	7.432* (2.10)
NEI	0.000895*** (24.20)	0.00143 (1.49)	0.000873*** (29.52)	0.00241 (1.72)
OPI	-0.000430*** (-15.56)	0.000252 (0.40)	-0.000411*** (-17.94)	-0.00152 (-1.53)
PRI	-0.000903 (-0.78)	0.0814** (2.62)	0.00110 (1.24)	-0.0364 (-0.85)
SAL	0.0000254*** (3.72)	- (-0.98)	0.0000175*** (3.33)	0.000652** (2.92)
Constant	27.95*** (31.38)	42.42 (1.76)	25.33*** (69.92)	29.57 (1.80)
Observations	5490	5345	11873	11660

Given that the GMM standard errors are downward biased, robust standard errors are recommended. A robust version of the Sargan test however is available in STATA after specifying `vce(robust)`. Given the limitations associated with the relatively short time span covered in our panel data set, we do not include any time trend component. Tests of joint significance are conducted but not reported. In line with the arbitrary rule of thumb suggested by Roodman (2009), the number of instruments doesn't the individual units (number of groups) in the panel suggesting potential problems of instrument proliferation are not apparent.

INT is found statistically insignificant with ROE for both proxies INT and INT/AST. With ROA, it is insignificant for INT but significant for INT/AST. Mostly, this result supports the proposition that intangible assets do not have significant effect on firm's financial performance. However, inconsistency of the results indicates the heterogeneity among the firms in term of asset class, nature of sectors. It also supports the possibility of multiple regimes in focus variables.

Sectoral Analysis

While aggregate analysis of financial performance offers an inclusive understanding of the effects intangible assets, a further sectoral analysis is able to give a more comprehensive understanding on the issue according to each industrial sector. There is a possibility of sector-specific growth, hence, there is a need to look at sector-specific sensitiveness (Sehrawat & Giri, 2017). A large collection of heterogenous firms may also introduce statistical regularities that are only the result of the aggregation procedure (e.g., via Central Limit Theorem); however, such aggregate analyses may lead to ambiguous conclusions (Bottazzi, & Secchi, 2003).

The sectoral differences in the coupling of revenues to outputs also imply greater pressure to improve performance in for-profit sectors (Kalleberg, Marsden, Reynolds, & Knoke, 2006). It is important to note that the stride for profit and related activities fluctuate for sectors; for instance, financial, materials, and telecommunication service sectors are more volatile than healthcare, energy and consumer staples sectors (Bottazzi & Secchi, 2003). Innovation activities in some service sectors such as telecommunications, transports and finance are associated with the establishment of expensive technological infrastructures, which require large financial resources and high demand. Consequently, for firms in these sectors, past economic performances might be more relevant as a basis for their overall financial commitment to innovation (Cainelli, Evangelista, & Savona, 2005). Hence, there is the existence of widespread heterogeneity within each class and within each sector, as the production processes in quite diverse ways, and such heterogeneity does not occur with the same characteristics across industries (Bottazzi, Secchi, & Tamagni, 2007). In reality, sectoral differences in dividend yields, capitalisations, and number of firms admitted to the sector accounted for more than two-third of the changes in market share. (Siegel & Schwartz, 2006). Therefore, for the existence of sectoral specificities in business operation, the 'pooling' of firms operating in different industrial sectors may conceal the specific characteristics of the dynamics of firms operating in different sectors (Bottazzi & Secchi, 2003). This calls for the need for data disaggregation to make more meaningful analysis.

The whole data samples are splitted into 10 sectors, followed by S&P methodology.

TABLE 5
SECTORAL CODE

Sector Name	Code
Consumer Discretionary	1
Consumer Staples	2
Energy	3
Financials	4
Health Care	5
Industrials	6
Information Technology	7
Materials	8
Telecommunication Services	9
Utilities	10

TABLE 6
SECTOR: CONSUMER DISCRETIONARY

	ROA	ROE	ROA	ROE
	INT	INT	INT/AST	INT/AST
L.ROA	0.0201		0.0350	
	(1.28)		(1.43)	
L.ROE		0.0705		0.122**
		(0.95)		(2.63)
INT	0.0000202	-0.00206		
	(0.51)	(-1.82)		
INTAST			0.271	-12.24*
			(0.52)	(-2.09)
LAST	-4.029***	-42.45**	-3.137***	-12.84***
	(-8.49)	(-3.26)	(-16.00)	(-6.44)
LNEI	0.109	19.81**	1.535***	12.30***
	(0.44)	(2.58)	(5.78)	(3.96)
LEBI	1.162**	-27.16*	1.777***	0.408
	(2.99)	(-2.37)	(5.61)	(0.11)
NEI	0.00728***	-0.0197	0.00210*	-0.0169
	(9.83)	(-0.87)	(2.38)	(-1.65)
OPI	-0.00133	0.0304	-0.00183**	0.00866
	(-1.76)	(1.44)	(-2.75)	(1.14)
PRI	0.000101	0.114	0.00257	0.0333
	(0.04)	(1.59)	(0.60)	(0.75)
SAL	-0.0000517	0.00176	0.0000884	0.000225
	(-0.62)	(0.77)	(1.31)	(0.30)
Constant	31.48***	444.0***	12.18***	58.66***
	(7.56)	(3.85)	(17.92)	(8.59)
Observations	25	25	66	66

Given that the GMM standard errors are downward biased, robust standard errors are recommended. A robust version of the Sargan test however is available in STATA after specifying `vce(robust)`. Given the limitations associated with the relatively short time span covered in our panel data set, we do not include any time trend component. Tests of joint significance are conducted but not reported. In line with the arbitrary rule of thumb suggested by Roodman (2009), the number of instruments doesn't the individual units (number of groups) in the panel suggesting potential problems of instrument proliferation are not apparent.

In table 6, For intangible assets, we find that it is statistically insignificant with ROA and ROE, but significant relationship is found between ROE and INTAST and it is negatively correlated.

TABLE 7
SECTOR: CONSUMER STAPLES

	ROA	ROE	ROA	ROE
	INT	INT	INT/AST	INT/AST
L.ROA	0.134*** (5.22)		0.125*** (6.18)	
L.ROE		0.574*** (10.31)		-0.0285 (-0.63)
INT	-0.0000424 (-1.24)	- (-1.39)		
INTAST			-1.215 (-1.04)	-714.5 (-1.28)
LAST	-5.983*** (-16.02)	-13.72* (-2.07)	-4.432*** (-21.26)	-127.2 (-1.39)
LNEI	-0.555* (-2.27)	-5.541 (-1.18)	1.251*** (9.47)	12.97 (0.22)
LEBI	6.521*** (11.75)	15.38 (1.48)	2.591*** (13.05)	35.49 (0.40)
NEI	0.00129*** (6.26)	0.00737 (1.84)	0.000416 (1.92)	-0.0199 (-0.20)
OPI	-0.00118*** (-6.17)	-0.00359 (-0.92)	-0.000352* (-2.16)	-0.0425 (-0.55)
PRI	0.00596* (2.27)	0.0385 (0.88)	0.00750* (2.40)	0.0310 (0.02)
SAL	0.0000461** (3.08)	0.000318 (1.19)	0.0000349* (2.19)	0.0189** (2.67)
Constant	19.51*** (10.36)	59.40 (1.73)	20.10*** (14.76)	675.6 (1.12)
Observations	137	135	297	295

Given that the GMM standard errors are downward biased, robust standard errors are recommended. A robust version of the Sargan test however is available in STATA after specifying `vce(robust)`. Given the limitations associated with the relatively short time span covered in our panel data set, we do not include any time trend component. Tests of joint significance are conducted but not reported. In line with the arbitrary rule of thumb suggested by Roodman (2009), the number of instruments doesn't the individual units (number of groups) in the panel suggesting potential problems of instrument proliferation are not apparent.

In consumer staple sector table 7, both of the proxies of intangible assets are found statistically insignificant to financial performance. Results reveal that, intangible assets do not have any impact on firms' financial performance.

TABLE 8
SECTOR: ENERGY

	ROA	ROE	ROA	ROE
	INT	INT	INT/AST	INT/AST
L.ROA	-0.0178 (.)		-0.0332 (-1.08)	
L.ROE		0.0117 (.)		0.0664 (1.64)
INT	-0.00322 (.)	-0.00748 (.)		
INTAST			-9.512 (-1.42)	-50.05** (-3.17)
LAST			-4.686*** (-25.29)	-6.203*** (-14.84)
LNEI			2.115*** (3.56)	0.722 (0.52)
LEBI	5.631 (.)	14.32 (.)	3.129*** (4.81)	5.333*** (3.55)
NEI	0.00395 (.)	0.00533 (.)	0.00226*** (3.57)	0.00675*** (4.56)
OPI	0.00993 (.)	0.0340 (.)	-0.00206*** (-3.83)	-0.00310* (-2.53)
PRI	-0.890 (.)	-2.693 (.)	0.00561 (0.55)	0.0404 (1.68)
SAL	0.0000120 (.)	0.0000523 (.)	0.00000879 (1.44)	- (-0.38)
Constant	-37.40 (.)	-100.1 (.)	16.06*** (11.45)	26.31*** (7.98)
Observations	8	8	36	36

Given that the GMM standard errors are downward biased, robust standard errors are recommended. A robust version of the Sargan test however is available in STATA after specifying `vce(robust)`. Given the limitations associated with the relatively short time span covered in our panel data set, we do not include any time trend component. Tests of joint significance are conducted but not reported. In line with the arbitrary rule of thumb suggested by Roodman (2009), the number of instruments doesn't the individual units (number of groups) in the panel suggesting potential problems of instrument proliferation are not apparent.

In energy sector table 8, both of the proxies of intangible assets are found statistically insignificant to financial performance. Results reveal that, intangible assets do not have any impact on firms' financial performance.

TABLE 9
SECTOR: FINANCIALS

	ROA	ROE	ROA	ROE
	INT	INT	INT/AST	INT/AST
L.ROA	0.0932 [*]		0.124 ^{***}	
	(2.20)		(3.59)	
L.ROE		0.997 ^{***}		1.061 ^{***}
		(11.89)		(18.60)
INT	-0.000779	0.00666		
	(-1.40)	(0.90)		
INTAST			6.510	118.8
			(0.62)	(1.16)
	(-11.11)	(-4.61)	(-16.57)	(-2.80)
LNEI	0.378	-2.497	0.388	-4.553 [*]
	(0.73)	(-0.43)	(1.37)	(-2.17)
LEBI	8.429 ^{***}	22.18 [*]	7.036 ^{***}	11.95 ^{**}
	(8.13)	(2.12)	(12.67)	(2.64)
NEI	0.00242 ^{***}	0.0130 ^{**}	0.00277 ^{***}	0.0142 ^{***}
	(5.77)	(2.92)	(9.28)	(6.33)
OPI	-0.00199 ^{***}	-0.00248	-0.00184 ^{***}	-0.00156
	(-6.22)	(-0.75)	(-7.83)	(-0.83)
PRI	0.0139	0.0883	0.0209 ^{**}	-0.0621
	(1.93)	(1.16)	(3.22)	(-1.13)
SAL	0.000113 ^{**}	0.00104 ^{**}	0.0000925 ^{**}	0.0000312
	(3.27)	(2.84)	(3.08)	(0.13)
Constant	23.66 ^{***}	135.0 ^{***}	26.60 ^{***}	28.54 ^{**}
	(9.54)	(5.17)	(19.54)	(2.94)
Observations	55	55	100	100

Given that the GMM standard errors are downward biased, robust standard errors are recommended. A robust version of the Sargan test however is available in STATA after specifying `vce(robust)`. Given the limitations associated with the relatively short time span covered in our panel data set, we do not include any time trend component. Tests of joint significance are conducted but not reported. In line with the arbitrary rule of thumb suggested by Roodman (2009), the number of instruments doesn't the individual units (number of groups) in the panel suggesting potential problems of instrument proliferation are not apparent.

In financial sector, table 9, intangible assets is found insignificant in financial sector, that means, intangible assets do not have any effect on firms' financial performance financial sector. Although this result can be different with different terms.

TABLE 10
SECTOR: HEALTH CARE

	ROA	ROE	ROA	ROE
	INT	INT	INT/AST	INT/AST
L.ROA	-0.0869*	-0.0869*	-0.0818***	
	(-2.42)	(-2.42)	(-5.51)	
L.ROE				-0.0633***
				(-4.78)
INT	0.00143	0.00143		
	(1.29)	(1.29)		
INTAST			48.00***	191.2**
			(4.49)	(2.87)
LAST	-1.614	-1.614	-3.728***	-10.48***
	(-0.19)	(-0.19)	(-9.40)	(-7.54)
LNEI	1.632	1.632	0.484	2.229*
	(0.45)	(0.45)	(1.64)	(2.16)
LEBI	-2.523	-2.523	2.694***	6.370***
	(-0.41)	(-0.41)	(5.51)	(4.19)
NEI	0.00269**	0.00269**	0.00313***	0.00958***
	(2.85)	(2.85)	(23.77)	(6.81)
OPI	-0.00124	-0.00124	-0.00169***	-0.00345**
	(-0.94)	(-0.94)	(-5.67)	(-2.63)
PRI	0.138*	0.138*	0.0565***	0.217***
	(2.12)	(2.12)	(3.42)	(3.45)
SAL	-0.000107	-0.000107	-0.0000490	0.0000505
	(-0.37)	(-0.37)	(-0.76)	(0.17)
Constant	24.73	24.73	18.02***	50.88***
	(0.37)	(0.37)	(8.78)	(6.74)
Observations	27	27	98	86

Given that the GMM standard errors are downward biased, robust standard errors are recommended. A robust version of the Sargan test however is available in STATA after specifying `vce(robust)`. Given the limitations associated with the relatively short time span covered in our panel data set, we do not include any time trend component. Tests of joint significance are conducted but not reported. In line with the arbitrary rule of thumb suggested by Roodman (2009), the number of instruments doesn't the individual units (number of groups) in the panel suggesting potential problems of instrument proliferation are not apparent.

In Health care sector, table 10, unlike other sectors, intangible assets is found highly significant and positively correlated with both proxies of financial performance ROA and ROE. This results reveal that, higher expenditures in intangible assets help firms to perform better financially in healthcare sector.

The analysis failed to apply GMM technique to compute the values for industrials sectors because of inadequate data.

TABLE 11
SECTOR: INFORMATION TECHNOLOGY

	ROA	ROE	ROA	ROE
	INT	INT	INT/AST	INT/AST
L.ROA	-0.152 (.)		0.00125 (0.02)	
L.ROE		-0.0675 (.)		-0.0554 (-0.88)
INT	-0.000581 (.)	-0.000673 (.)		
INTAST			7.135 (0.26)	-97.53 (-1.57)
LAST	-6.426 (.)	-12.21 (.)	-0.261 (-0.16)	2.015 (0.57)
LNEI	0.754 (.)	2.164 (.)	0.636** (2.59)	1.594** (2.94)
LEBI	-1.664 (.)	-4.266 (.)	2.003* (2.09)	7.030** (3.27)
NEI	0.00306 (.)	0.00602 (.)	0.00386*** (4.24)	0.00968*** (4.79)
OPI	0.00000795 (.)	0.00109 (.)	-0.00128 (-1.48)	-0.00490* (-2.51)
PRI	-0.0210 (.)	-0.0415 (.)	-0.0284 (-1.19)	-0.0297 (-0.55)
SAL	0.0000644 (.)	- 0.0000404 (.)	-0.000212 (-1.23)	-0.000512 (-1.35)
Constant	74.36 (.)	146.4 (.)	-7.592 (-0.57)	-54.93 (-1.90)
Observations	11	11	29	29

Given that the GMM standard errors are downward biased, robust standard errors are recommended. A robust version of the Sargan test however is available in STATA after specifying `vce(robust)`. Given the limitations associated with the relatively short time span covered in our panel data set, we do not include any time trend component. Tests of joint significance are conducted but not reported. In line with the arbitrary rule of thumb suggested by Roodman (2009), the number of instruments doesn't the individual units (number of groups) in the panel suggesting potential problems of instrument proliferation are not apparent.

Intangible assets are not found to have any impact on financial performance of the firms.

TABLE 12
SECTOR: MATERIALS

	ROA	ROE	ROA	ROE
	INT	INT	INT/AST	INT/AST
L.ROA	0.0583 (1.53)		0.0134 (0.87)	
L.ROE		7.983*** (9.65)		0.865*** (4.02)
INT	-0.00182 (-1.76)	0.0404 (0.73)		
INTAST			-2.376 (-0.51) (-24.91)	971.8*** (6.76) (-0.21)
LNEI	1.628 (1.57)	346.3*** (5.55)	0.532 (1.40)	-1.766 (-0.14)
LEBI	4.216*** (3.55)	-437.5*** (-5.75)	7.592*** (14.40)	-0.0188 (-0.00)
NEI	0.00885*** (4.89)	-0.791*** (-7.31)	0.0195*** (29.81)	0.0282 (1.30)
OPI	-0.00245* (-2.41)	0.505*** (8.00)	-0.00888*** (-15.02)	0.0355 (1.81)
PRI	0.00571 (0.56)	1.577* (1.98)	-0.0252* (-2.33)	0.869* (2.24)
SAL	-0.000201** (-2.89)	-0.0186*** (-3.97)	-0.0000295 (-0.46)	-0.00865*** (-4.09)
Constant	20.05*** (10.40)	171.3 (1.31)	19.56*** (21.42)	29.07 (0.99)
Observations	67	64	178	175

Given that the GMM standard errors are downward biased, robust standard errors are recommended. A robust version of the Sargan test however is available in STATA after specifying `vce(robust)`. Given the limitations associated with the relatively short time span covered in our panel data set, we do not include any time trend component. Tests of joint significance are conducted but not reported. In line with the arbitrary rule of thumb suggested by Roodman (2009), the number of instruments doesn't the individual units (number of groups) in the panel suggesting potential problems of instrument proliferation are not apparent.

In material sector, table 12, intangible assets ratio is found significant and positively correlated with ROE, but insignificant for ROA.

TABLE 13
SECTOR: TELECOMMUNICATION SERVICES

	ROA	ROE	ROA	ROE
	INT	INT	INT/AST	INT/AST
L.ROA	0.0176 (0.33)		0.0704 (1.93)	
L.ROE		0.111 (1.81)		0.186*** (3.92)
INT	0.00131*** (3.45)	-0.00272 (-1.54)		
INTAST			0.0955 (0.05)	-6.295 (-0.63)
LNEI	(-11.44) -1.171 (-1.36)	(-3.34) -2.696 (-0.68)	(-15.60) -0.488 (-1.59)	(-2.97) 3.606* (2.28)
LEBI	12.80*** (8.49)	20.49** (3.01)	11.15*** (14.25)	8.472* (2.26)
NEI	0.0164*** (4.82)	0.0218 (1.37)	0.0173*** (7.26)	0.0128 (1.00)
OPI	-0.00165 (-0.52)	-0.0250 (-1.80)	-0.00489* (-2.21)	-0.0109 (-0.93)
PRI	0.0201* (2.34)	-0.0569 (-1.50)	0.0215** (3.05)	-0.0287 (-0.77)
SAL	-0.00209*** (-4.97)	0.00297 (1.74)	-0.00108*** (-3.88)	0.000263 (0.19)
Constant	25.41*** (7.44)	14.54 (0.99)	23.12*** (10.57)	23.97* (2.08)
Observations	81	81	128	128

Given that the GMM standard errors are downward biased, robust standard errors are recommended. A robust version of the Sargan test however is available in STATA after specifying `vce(robust)`. Given the limitations associated with the relatively short time span covered in our panel data set, we do not include any time trend component. Tests of joint significance are conducted but not reported. In line with the arbitrary rule of thumb suggested by Roodman (2009), the number of instruments doesn't the individual units (number of groups) in the panel suggesting potential problems of instrument proliferation are not apparent.

Unlike many other sectors, for telecommunication services, intangible assets is found positively significant. Though other proxies show insignificant result. Results reveal that higher expenditures in intangible assets of the firms in telecommunication services will create better financial performance.

TABLE 14
SECTOR: UTILITIES

	ROA	ROE	ROA	ROE
	INT	INT	INT/AST	INT/AST
L.ROA	0.135 ^{***} (3.48)		0.205 ^{***} (5.93)	
L.ROE		0.172 ^{***} (3.53)		0.144 ^{***} (3.90)
INT	-0.00437 [*] (-2.34)	-0.00718 [*] (-2.22)		
INTAST			-3.732 (-0.45)	-24.94 (-1.67)
LAST	-6.475 ^{***} (-6.99)	-12.09 ^{***} (-6.59)	-9.080 ^{***} (-17.00)	-14.15 ^{***} (-14.73)
LNEI	2.385 (1.77)	4.545 (1.82)	5.278 ^{***} (7.60)	7.331 ^{***} (5.89)
LEBI	2.696 [*] (2.22)	3.005 (1.26)	2.548 ^{**} (3.19)	7.438 ^{***} (4.77)
NEI	0.0159 ^{***} (3.93)	0.0107 (1.60)	0.0170 ^{***} (3.49)	0.0198 [*] (2.52)
OPI	-0.0102 ^{***} (-3.48)	-0.00893 (-1.84)	-0.0164 ^{***} (-3.93)	-0.0282 ^{***} (-4.04)
PRI	0.0200 (1.47)	0.000600 (0.03)	0.0331 (1.77)	0.0517 (1.58)
SAL	0.00110 (1.58)	0.00399 ^{**} (3.09)	0.00239 ^{**} (2.69)	0.00479 ^{**} (3.04)
Constant	29.80 ^{***} (6.15)	61.61 ^{***} (5.46)	34.28 ^{***} (16.72)	41.86 ^{***} (11.17)
Observations	58	57	104	102

Given that the GMM standard errors are downward biased, robust standard errors are recommended. A robust version of the Sargan test however is available in STATA after specifying `vce(robust)`. Given the limitations associated with the relatively short time span covered in our panel data set, we do not include any time trend component. Tests of joint significance are conducted but not reported. In line with the arbitrary rule of thumb suggested by Roodman (2009), the number of instruments doesn't the individual units (number of groups) in the panel suggesting potential problems of instrument proliferation are not apparent.

Statistical result in utilities sectors is found to be quite different from other sectors. It's the only sector where intangible assets is found negatively correlated with financial performance for both Proxies ROA and ROE. It refers that higher expenditures in intangible assets will result in lower financial performance. One obvious reason behind such findings is that utilities does not require much intellectual assets rather it is essential for people.

**TABLE 15
SUMMARY OF THE RESULTS**

Effect of Intangible assets on Financial Performance (ROA, ROE)	INT	ROA,	Not-significant	Diverse results, Intangible assets mostly do not effect financial performance, exception noticed.
		ROE	Not-significant	
	INT/AST	ROA	Significant, Negative	
		ROE	Not-significant	

**TABLE 16
SECTORAL DIFFERENCES AND INTANGIBLE ASSETS –FINANCIAL PERFORMANCE RELATIONSHIPS**

Sectors			Findings	Comments
Sector 1: Consumer Discretionary	INT	ROA,	Not-significant	Intangible assets itself does not found to have any significant effect on financial performance. However, the Intangible assets/ Total assets is found to have significant impact on financial performance of companies from different sectors. For Financials and IT sectors, it is insignificant, For healthcare, materials and telcom sectors it is positively significant and consumer, energy and utilities sectors it is negatively significant
		ROE	Not-significant	
	INT/AST	ROA	Not-significant	
		ROE	Significant, negative	
Sector 2: Consumer Staples	INT	ROA,	Not-significant	
		ROE	Not-significant	
	INT/AST	ROA	Not-significant	
		ROE	Not-significant	
Sector 3: Energy	INT	ROA,	Not-significant	
		ROE	Not-significant	
	INT/AST	ROA	Not-significant	
		ROE	Significant, negative	
Sector 4: Financials	INT	ROA,	Not-significant	
		ROE	Not-significant	
	INT/AST	ROA	Not-significant	
		ROE	Not-significant	
Sector 5: Health Care	INT	ROA,	Not-significant	
		ROE	Not-significant	
	INT/AST	ROA	Significant, Positive	
		ROE	Significant, Positive	
Sector 6: Industrials Sector 7: Information Technology	INT	ROA,	Not-significant	
		ROE	Not-significant	
	INT/AST	ROA	Not-significant	
		ROE	Not-significant	

Sector 8: Materials	INT	ROA,	Not-significant
		ROE	Not-significant
	INT/AST	ROA	Not-significant
		ROE	Significant, Positive
Sector 9 : Telecommunication Services	INT	ROA,	Significant, Positive
		ROE	Not-significant
	INT/AST	ROA	Not-significant
		ROE	Not-significant
Sector 10: Utilities	INT	ROA,	Significant, Negative
		ROE	Significant, Negative
	INT/AST	ROA	Not-significant
		ROE	Not-significant

The current study intended to understand how such financial performance is influenced by firms' intangible assets. To do so, the study considers the ratio of intangible assets to firm's total assets and intangibles assets to measure the relationship. By employing the data from S&P 500 companies over the period from 1979 to 2015, the study finds diverse outcomes concerning the relationship. To attain comprehensive understanding the investigation further expands to sectoral analyses. Intangible assets itself is not found to have any significant effect on financial performance of the firms, although the Intangible assets/ Total assets ratio is found to have significant impact on financial performance of companies. For healthcare, materials and telecom sectors, the effect is positive and significant, while for consumer, energy and utilities sectors it is significant but negative; besides there was no significant statistical relationship was found for financials and IT sectors.

The research findings possess significant policy implications for different class of stakeholders as intangible assets are intense concerns for various parties. For the investors, the findings provide insights in risk-return paradigm in the framework of investment risk intangible assets holding by the firms and their subsequent return. Again, the sectoral performance differences indicate that investors needs to be sector sensitive while analysing firm's investment on intangible assets. The findings are also expected to help the financial managers to forecast the future return of a firm and also to measure the riskiness of financing and investment activities. In addition, this study contains noteworthy insights for the policy makers, government agencies and regulatory bodies; returns generated through intangible assets are vital to decide on the benefits, subsidization, taxation policy and such. Further, it is expected to aid the policy makers to settle on which sectors are worthy to be prioritized and how much be supported to get the expected result in line with country's investment policy. Also, notably the study adds value to the academia by considering intangibles' influence on corporate performance which is not clear in the existing literature. Besides answering some unsettled research problems and adding knowledge to the growing body of literature in this filed, the study unveils further avenue of research for academics.

The study endeavoured comprehensive analyses and a fairly novel attempt to understand the nexus; nevertheless, it is not devoid of some limitations mostly owing to unavailability of adequate data. The dataset comprises only S&P 500 companies which are predominantly large companies based on developed economy (i.e. the USA), thus leads to lack of generalizability of the findings for the companies around the globe. Also, in some cases sufficient sectoral data were not available and levied restrictions on analyses. Likewise, the study did not take into consideration the institutional and governance variables. Since significant difference is found in asset-equity structure of the companies, further analyses with such variables could have been more insightful. Hence, future research may consider new dataset and

incorporate regional analyses by giving consideration for institutional and political variables which will stretch better generalizability of the research. Furthermore, study can be further extended by considering threshold and asset size effect for different sectors as the current finding is somewhat heterogeneous.

ENDNOTES

1. A data analysis software
2. Over-identification test
3. Auto-correlation test

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